

Bayesian inference

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Overview of the talk

1 Probabilistic modelling and representation of uncertainty

1.1 Bayesian paradigm

1.2 Hierarchical models

1.3 Frequentist versus Bayesian inference

2 Notes on Bayesian inference

2.1 Variational methods (ReML, EM, VB)

2.2 Family inference

2.3 Group-level model comparison

3 SPM applications

3.1 aMRI segmentation

3.2 Decoding of brain images

3.3 Model-based fMRI analysis (with spatial priors)

3.4 Dynamic causal modelling

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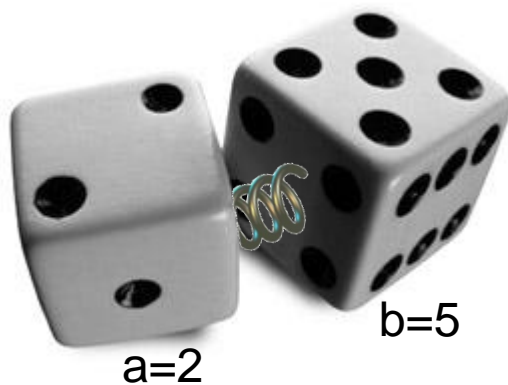
3.4 Dynamic causal modelling

Bayesian paradigm

probability theory: basics

Degree of *plausibility* desiderata:

- should be represented using real numbers (D1)
- should conform with intuition (D2)
- should be consistent (D3)



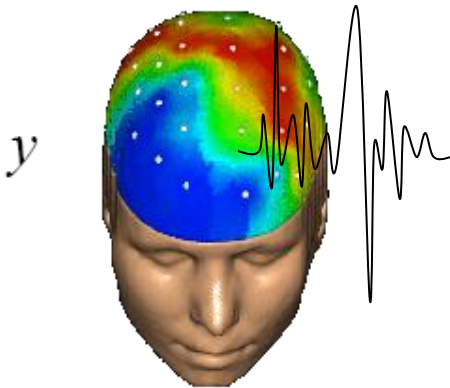
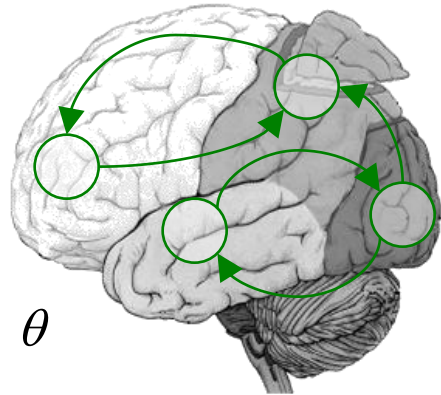
- normalization:
$$\sum_a P(a) = 1$$

- marginalization:
$$P(b) = \sum_a P(a, b)$$

- **conditioning :**
(Bayes rule)
$$P(a, b) = P(a|b)P(b)$$
$$= P(b|a)P(a)$$

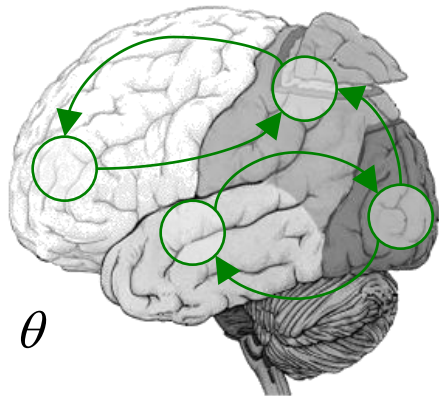
Bayesian paradigm

deriving the likelihood function

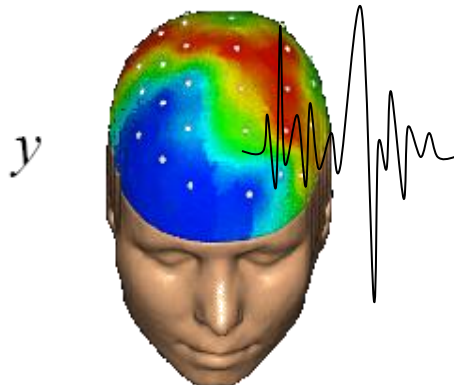


Bayesian paradigm

likelihood, priors and the model evidence



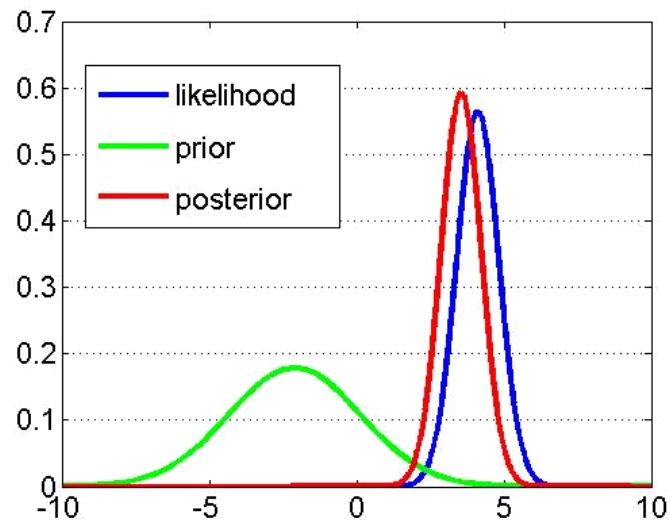
generative model m



Likelihood: $p(y|\theta, m)$

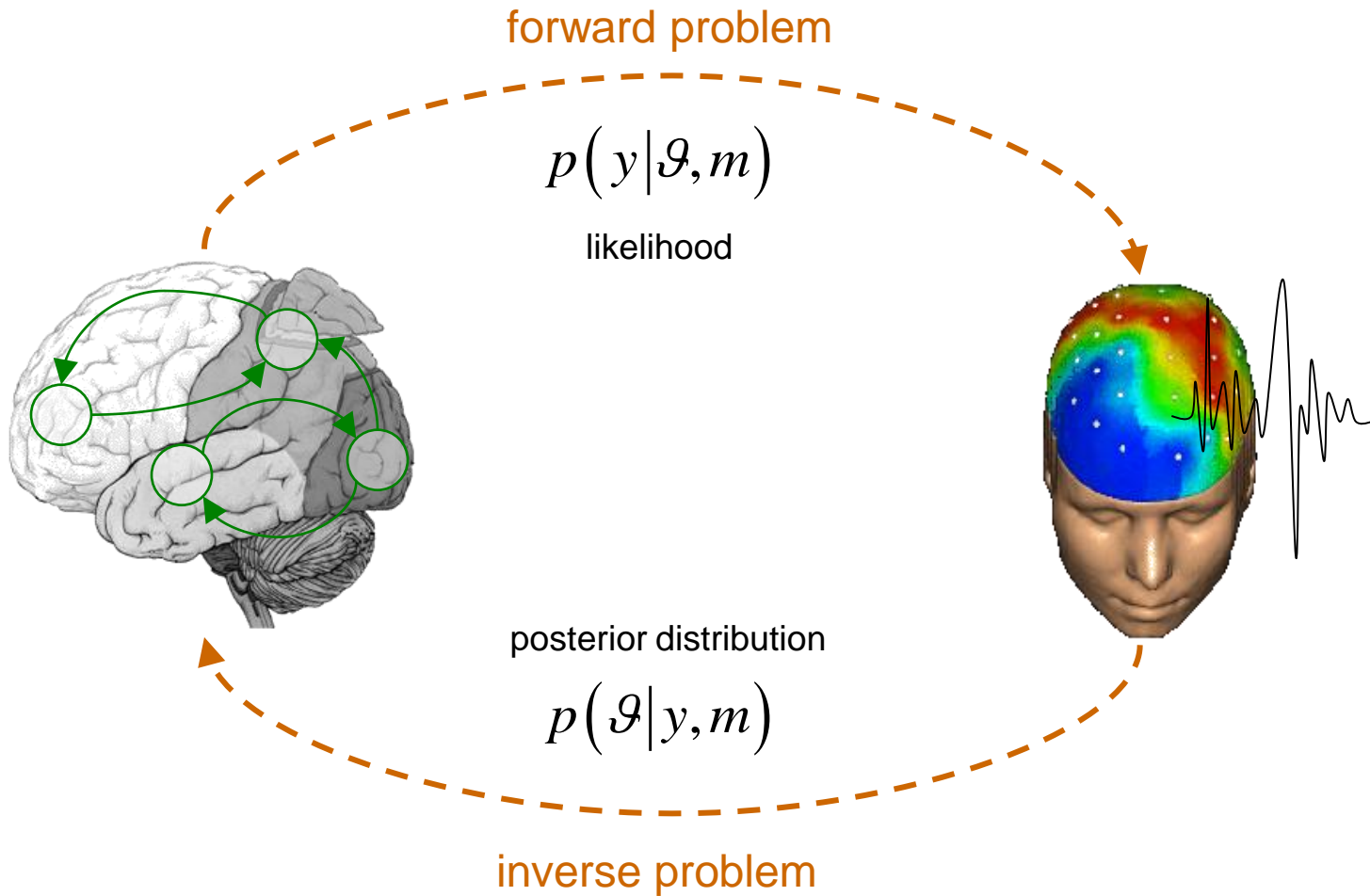
Prior: $p(\theta|m)$

Bayes rule: $p(\theta|y, m) = \frac{p(y|\theta, m) p(\theta|m)}{p(y|m)}$



Bayesian paradigm

forward and inverse problems

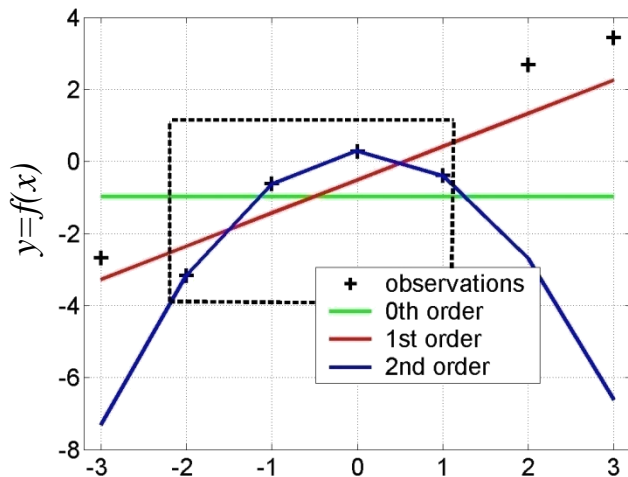
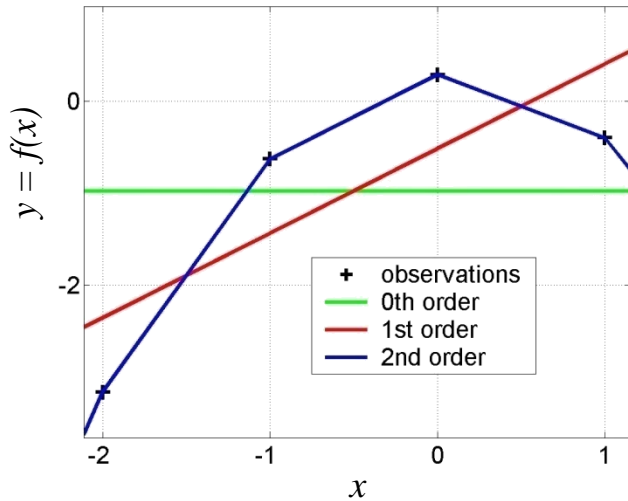


Bayesian paradigm

model comparison

Principle of parsimony :

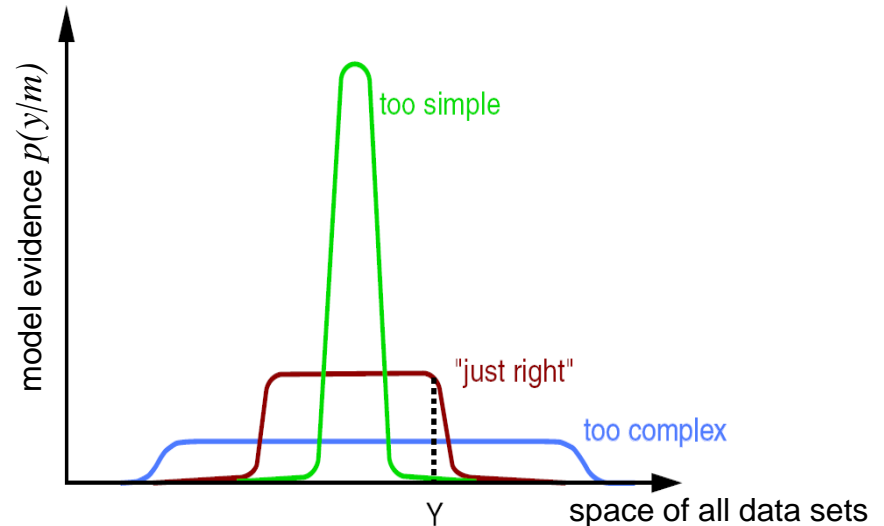
« plurality should not be assumed without necessity »



Model evidence:

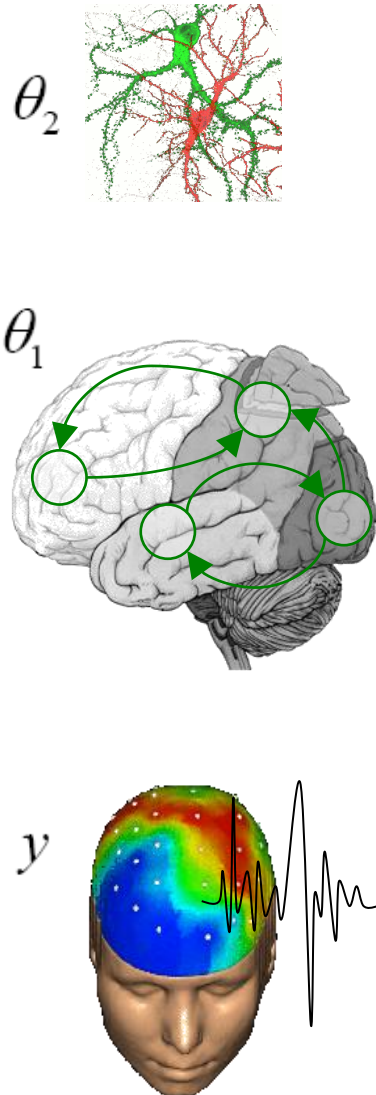
$$p(y|m) = \int p(y|\theta, m) p(\theta|m) d\theta$$

“Occam’s razor” :



Hierarchical models

principle



\vdots
 $p(\theta_2 | \theta_3, m)$

$$p(\theta_1 | \theta_2, m)$$

$$p(y | \theta_1, m)$$

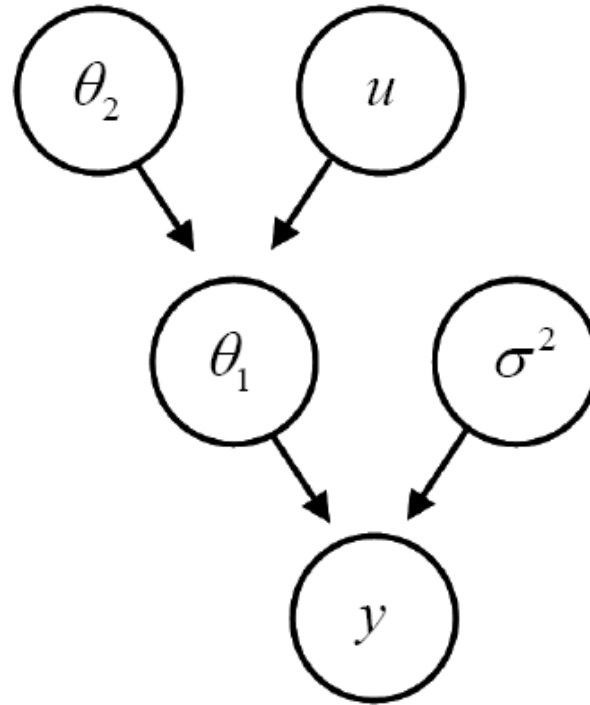
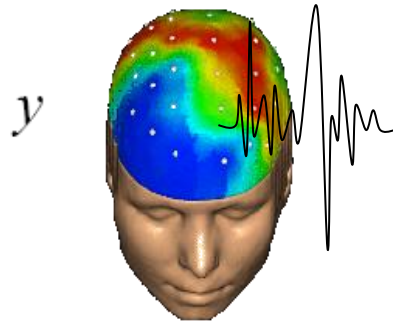
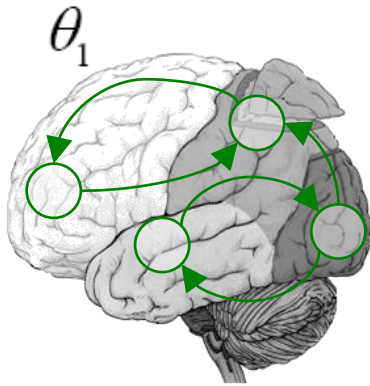
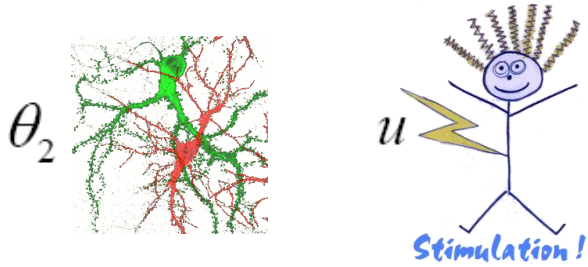
inference



causality

Hierarchical models

directed acyclic graphs (DAGs)



$$p(\theta_1 | \theta_2, u, m)$$

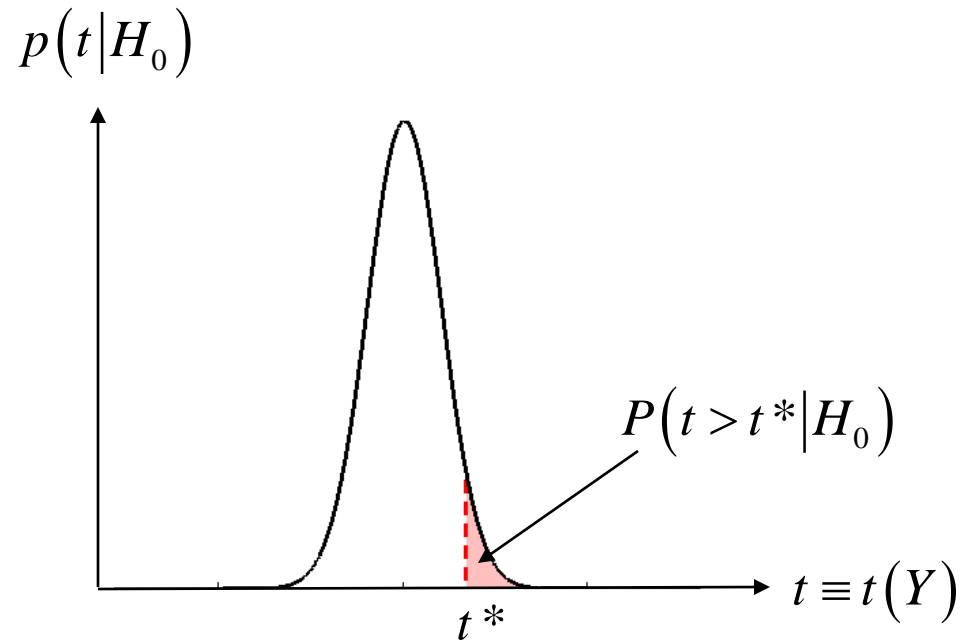
$$p(y | \theta_1, \sigma^2, m)$$

$$p(\theta | m) = \prod_j p(\theta_j | \text{par}(\theta_j), m)$$

Frequentist versus Bayesian inference

a (quick) note on hypothesis testing

- define the null, e.g.: $H_0 : \theta = 0$



- estimate parameters (obtain test stat.)
- apply decision rule, i.e.:
if $P(t > t^* | H_0) \leq \alpha$ then reject H_0

classical (null) hypothesis testing

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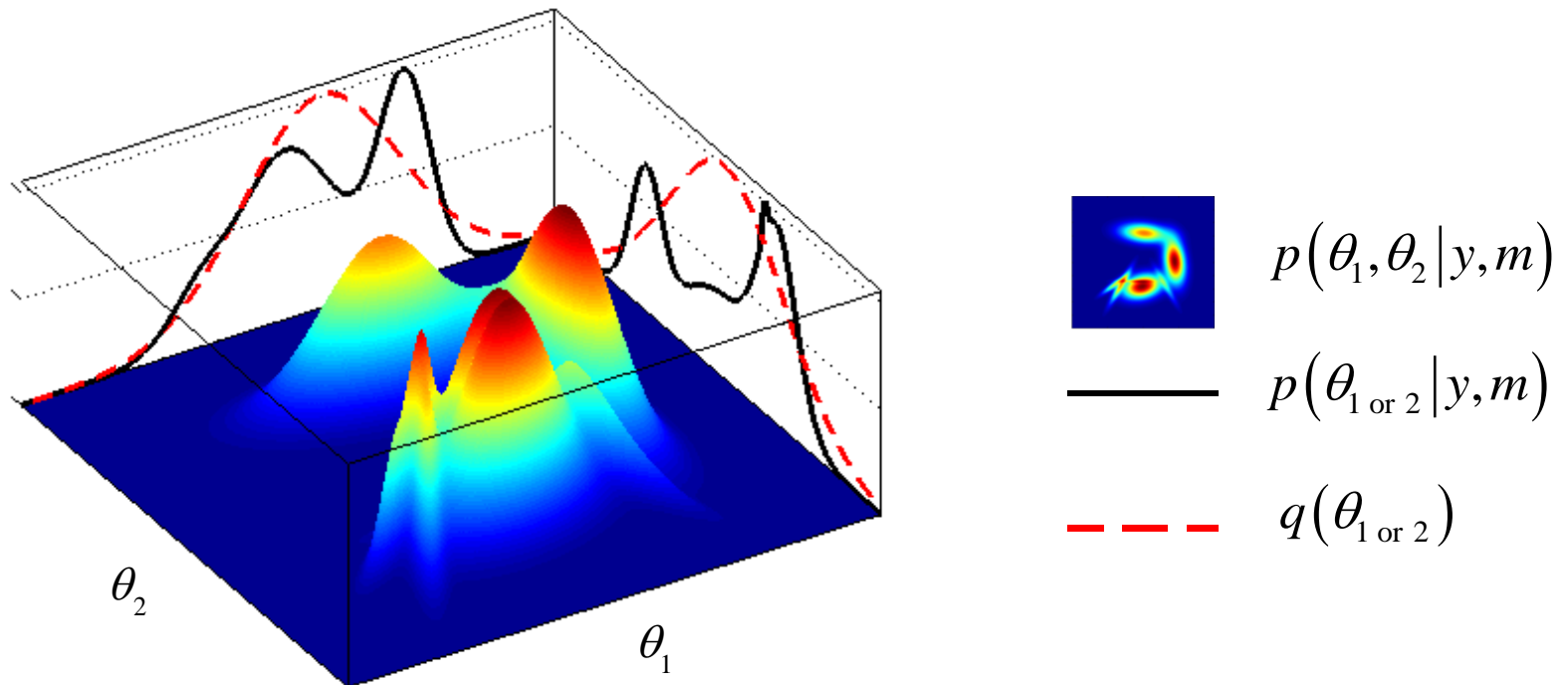
3.4 Dynamic causal modelling

Variational methods

VB / EM / ReML

$$\ln p(y|m) = \underbrace{\langle \ln p(\theta, y|m) \rangle_q + S(q)}_{\text{free energy } F(q)} + D_{KL}(q(\theta); p(\theta|y, m))$$

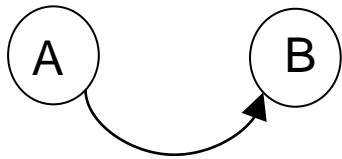
→ **VB** : maximize the **free energy** $F(q)$ w.r.t. the **approximate posterior** $q(\theta)$ under some (e.g., *mean field, Laplace*) simplifying constraint



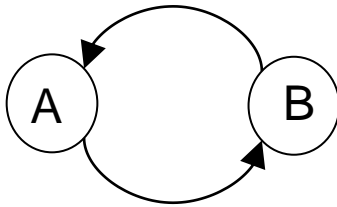
Family-level inference

trading inference resolution against statistical power

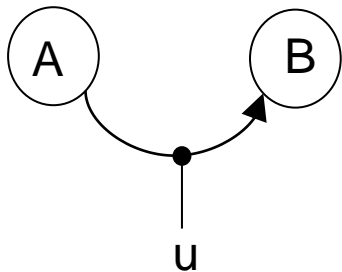
$$P(m_1|y) = 0.04$$



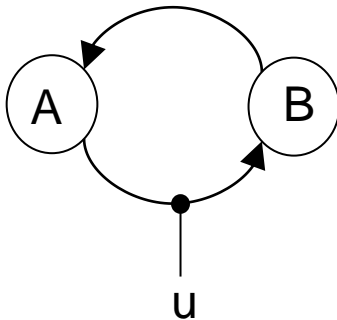
$$P(m_2|y) = 0.25$$



$$P(m_2|y) = 0.01$$



$$P(m_2|y) = 0.7$$

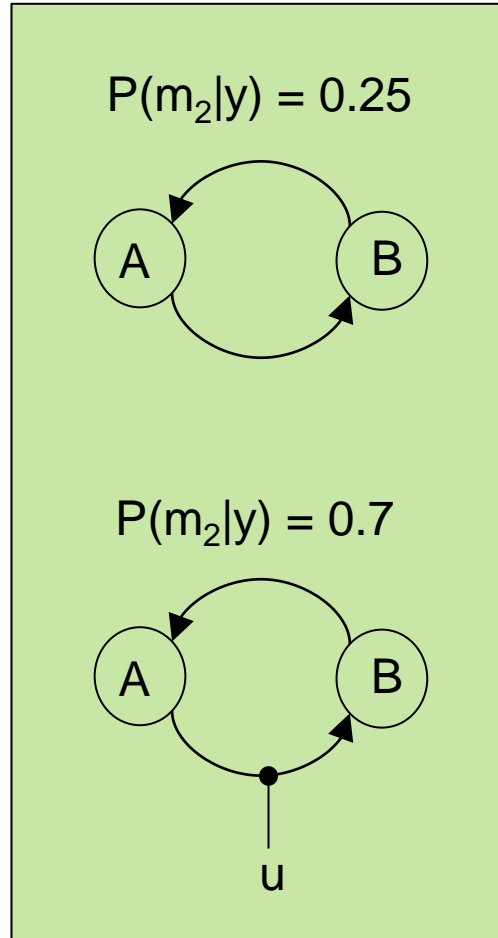
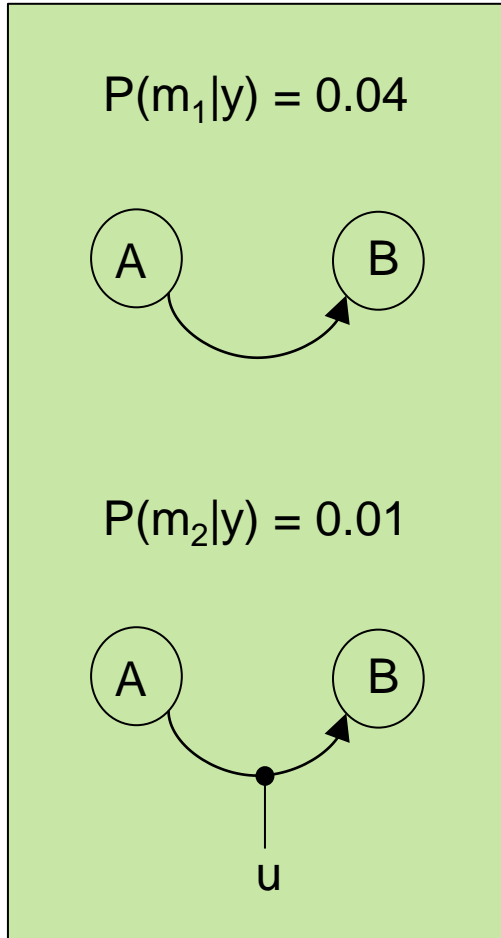


model selection error risk:

$$P(e = 1|y) = 1 - \max_m P(m|y) \\ = 0.3$$

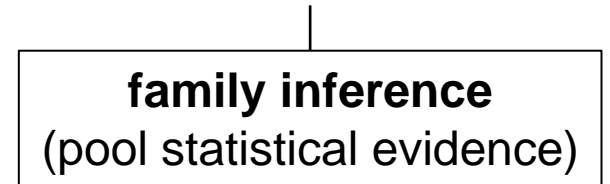
Family-level inference

trading inference resolution against statistical power



model selection error risk:

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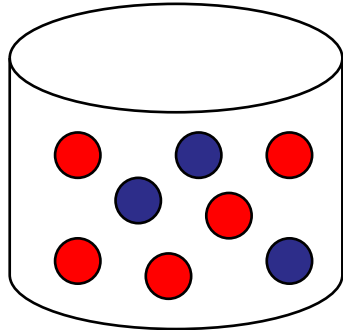


$$P(f|y) = \sum_{m \in f} P(m|y)$$

$$P(e = 1|y) = 1 - \max_f P(f|y) = 0.05$$

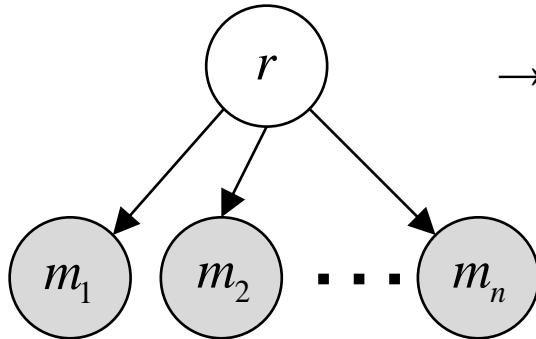
Group-level model comparison

preliminary: Polya's urn



$$\begin{cases} m_i = 1 & \rightarrow i^{\text{th}} \text{ marble is blue} \\ m_i = 0 & \rightarrow i^{\text{th}} \text{ marble is purple} \end{cases}$$

r = proportion of blue marbles in the urn



→ (binomial) probability of drawing a set of n marbles:

$$p(m|r) = \prod_{i=1}^n r^{m_i} (1-r)^{1-m_i}$$

Thus, our belief about the proportion of blue marbles is:

$$p(r|m) \propto p(r) \prod_{i=1}^n r^{m_i} (1-r)^{1-m_i} \stackrel{p(r) \propto 1}{\Rightarrow} E[r|m] = \frac{1}{n} \sum_{i=1}^n m_i$$

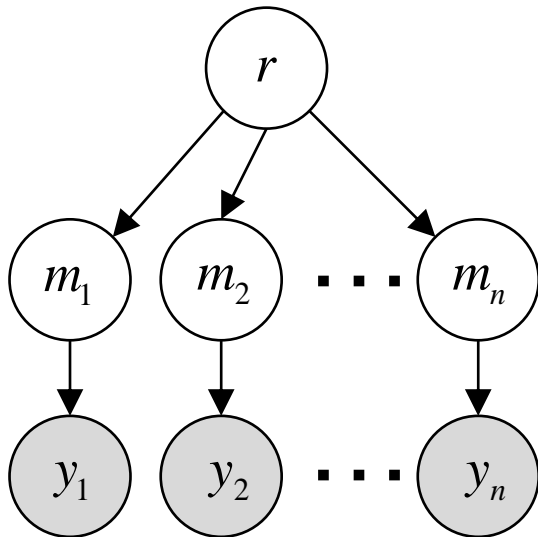
Group-level model comparison

what if we are colour blind?

At least, we can measure how likely is the i^{th} subject's data under each model!

● ● ■ ■ ■ ● ■ ■ ■ ●

$$p(y_1|m_1) \quad p(y_2|m_2) \quad \dots \quad p(y_i|m_i) \quad \dots \quad p(y_n|m_n)$$



$$p(r, m|y) \propto p(r) \prod_{i=1}^n p(y_i|m_i) p(m_i|r)$$

Our belief about the proportion of models is:

$$p(r|y) = \sum_m p(r, m|y)$$

Exceedance probability: $\varphi_k = P(r_k > r_{k' \neq k} | y)$

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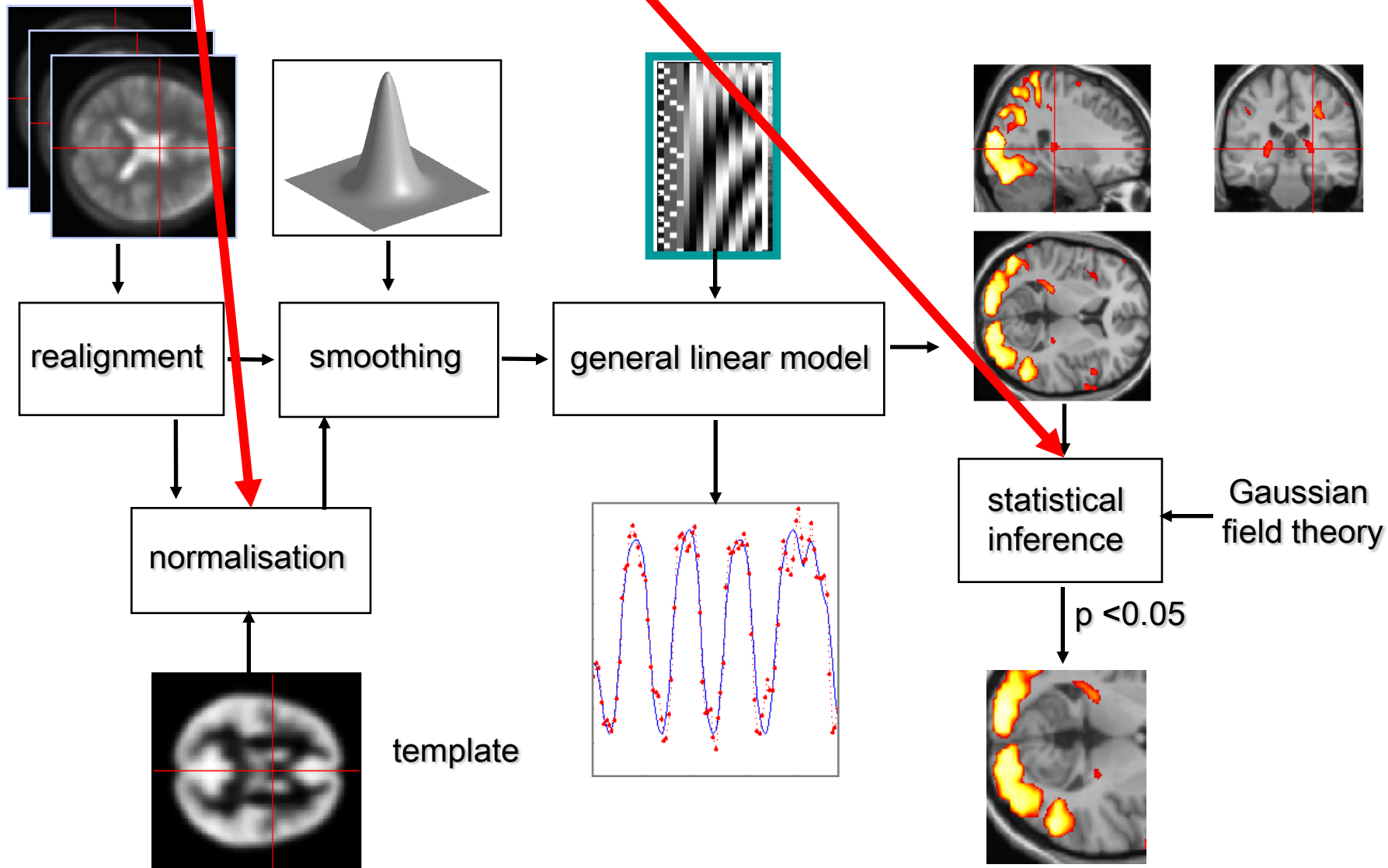
3.4 Dynamic causal modelling

segmentation and normalisation

posterior probability maps (PPMs)

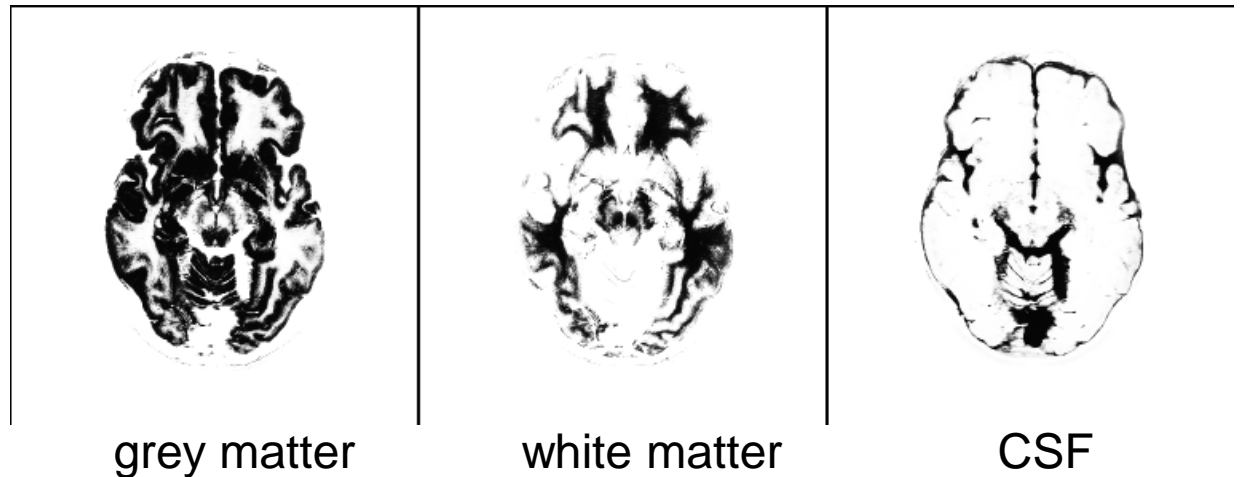
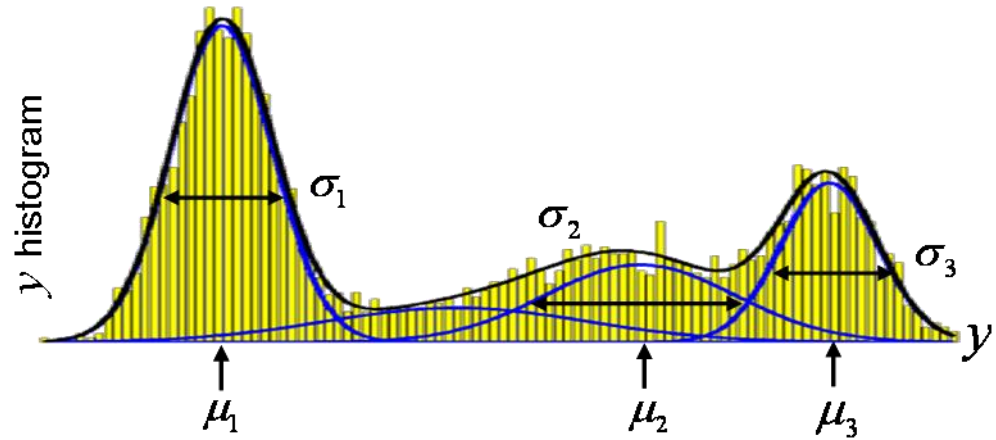
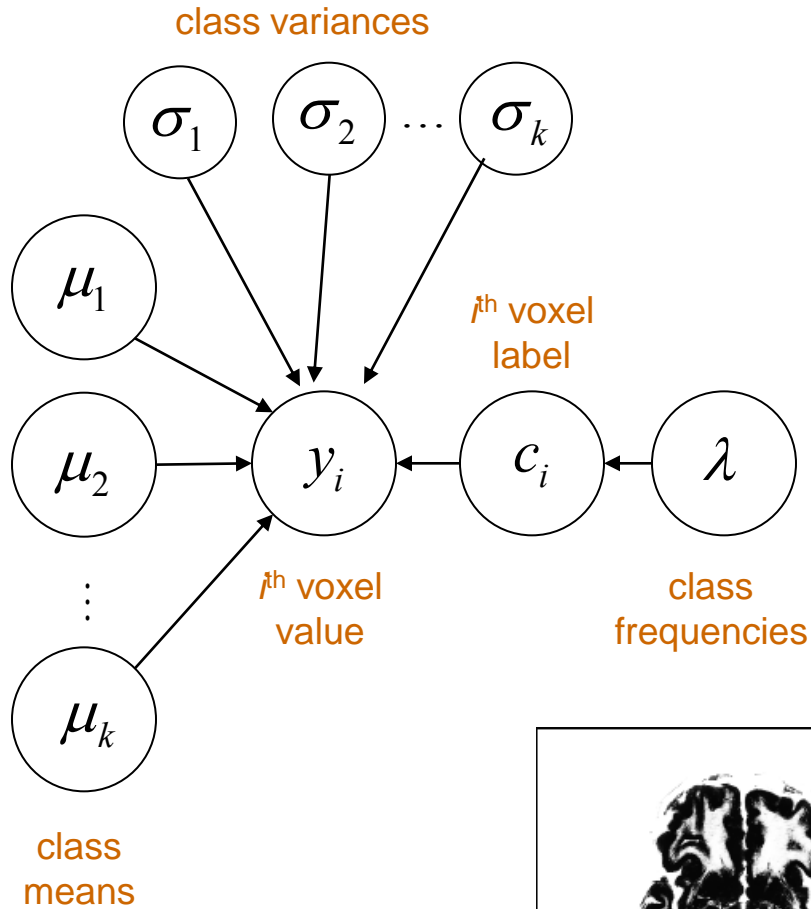
dynamic causal modelling

multivariate decoding



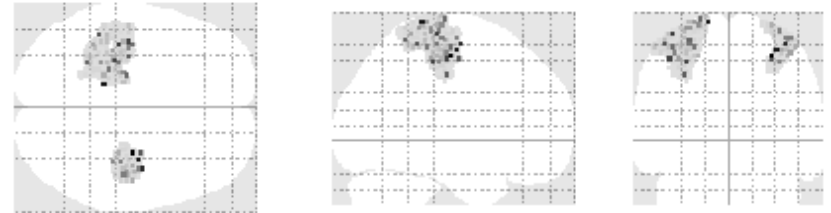
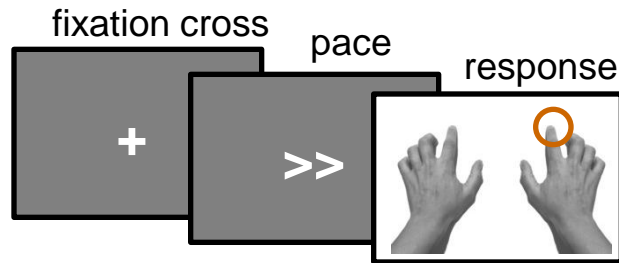
aMRI segmentation

mixture of Gaussians (MoG) model

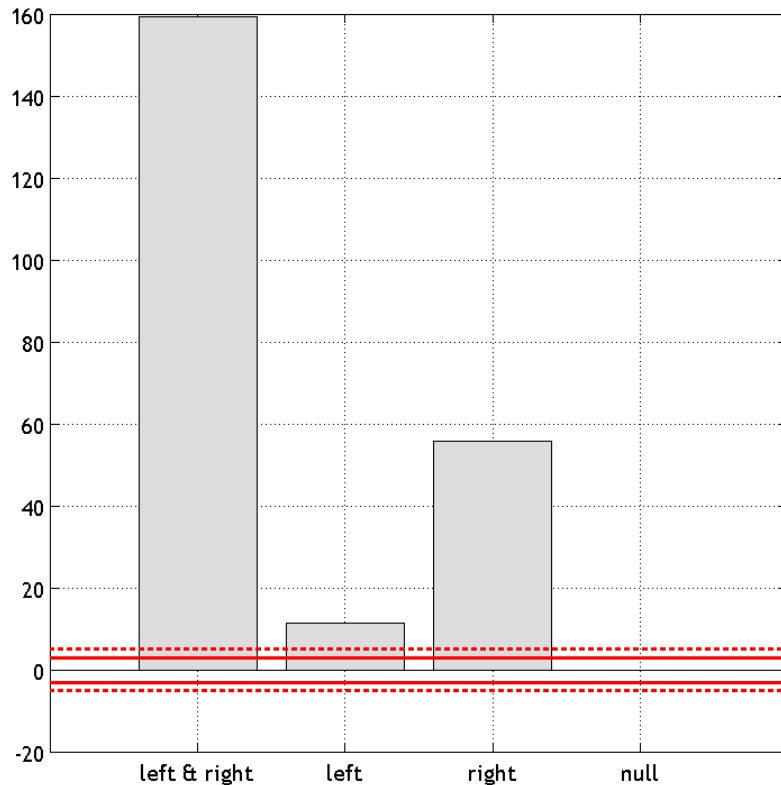


Decoding of brain images

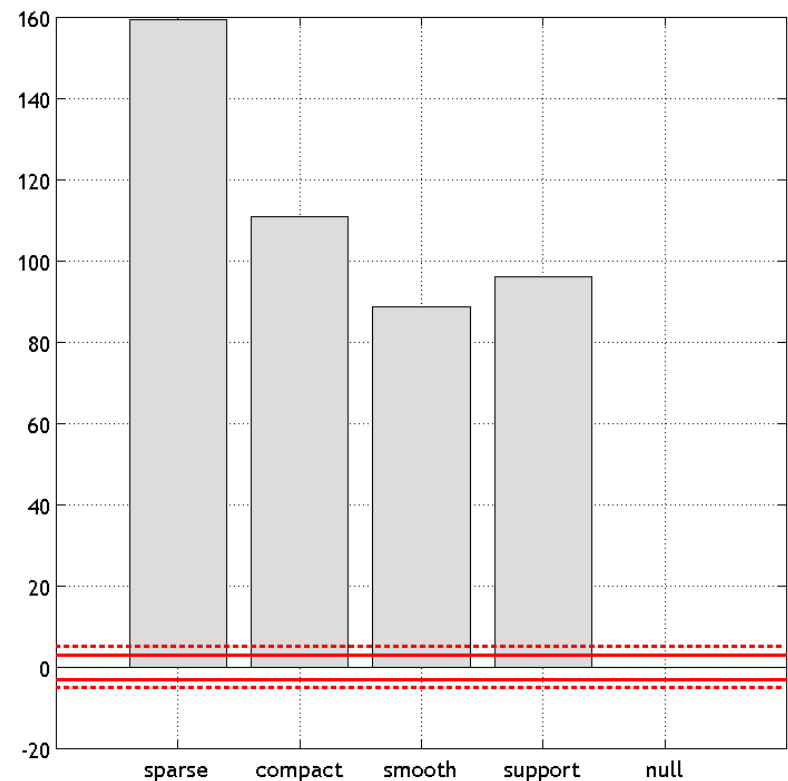
recognizing brain states from fMRI



log-evidence of X-Y sparse mappings:
effect of lateralization

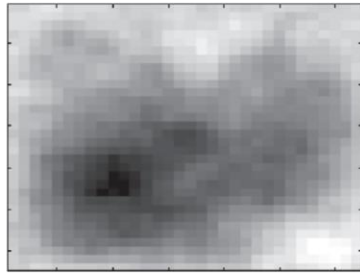
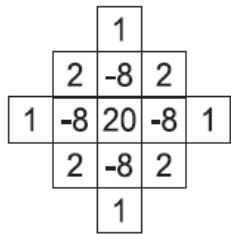


log-evidence of X-Y bilateral mappings:
effect of spatial deployment

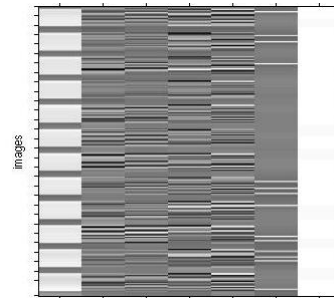


fMRI time series analysis

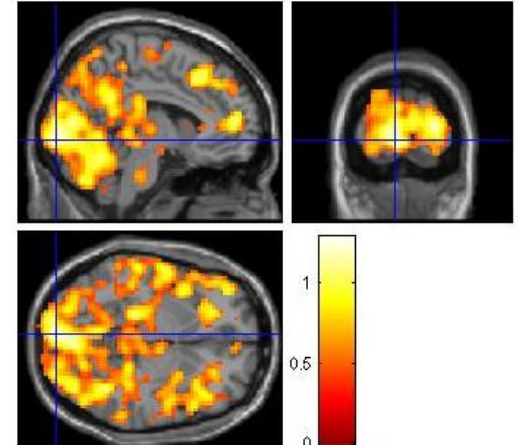
spatial priors and model comparison



short-term memory design matrix (X)



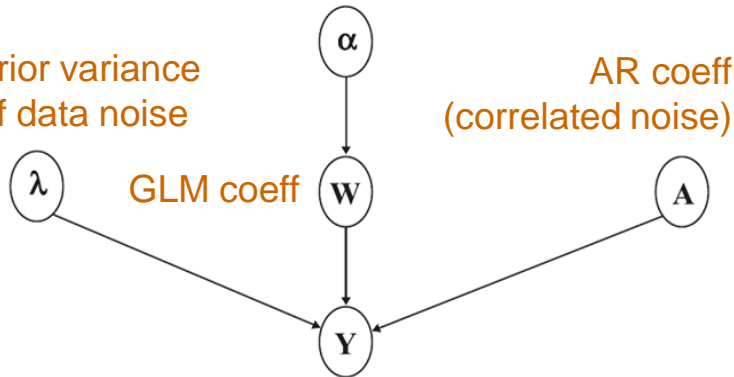
PPM: regions best explained by short-term memory model



prior variance of GLM coeff

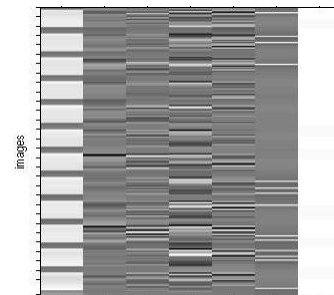
prior variance of data noise

AR coeff (correlated noise)

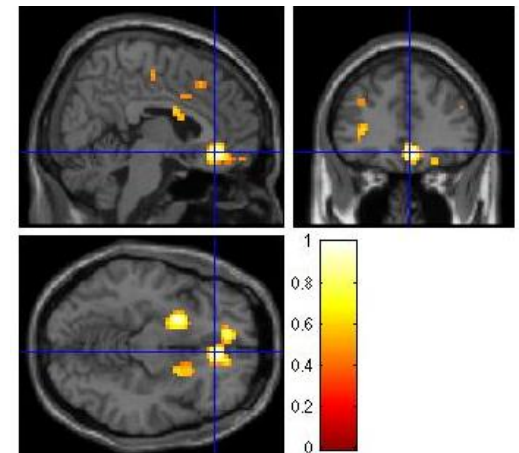


fMRI time series

long-term memory design matrix (X)

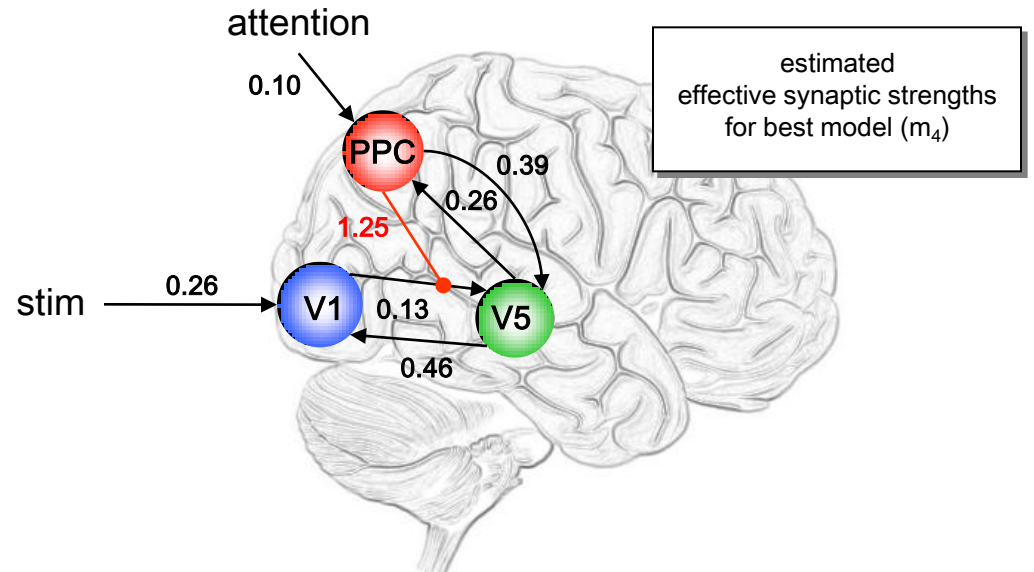
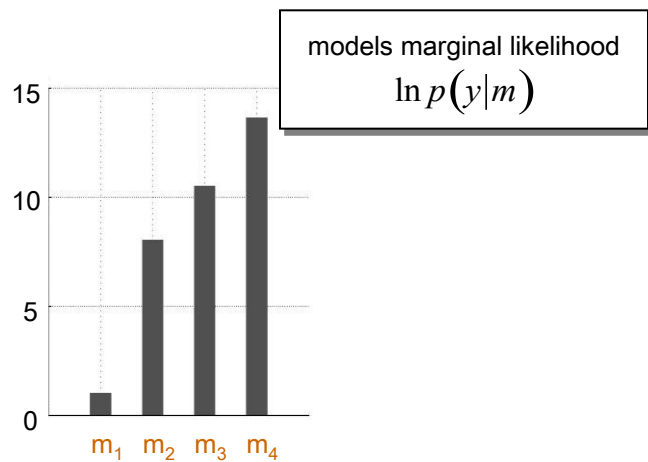
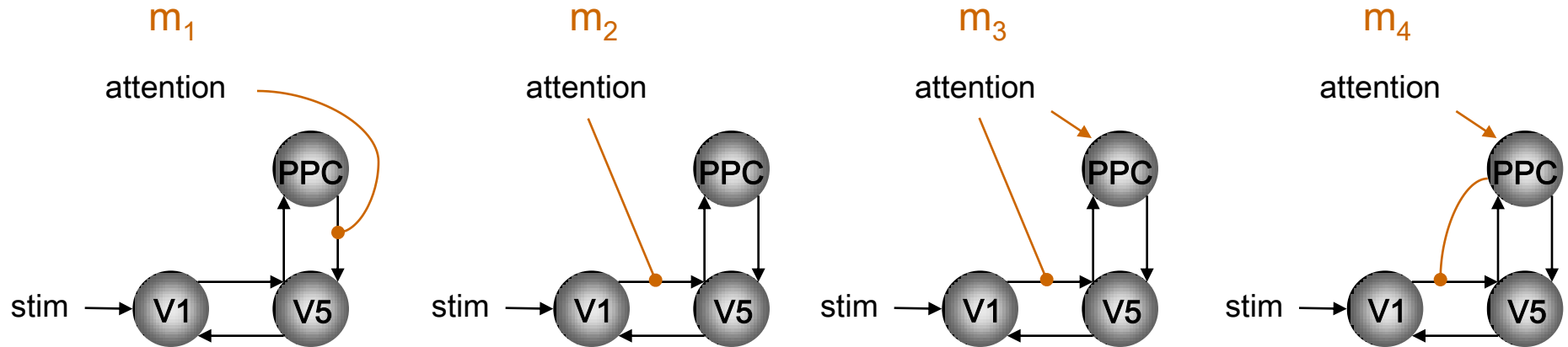


PPM: regions best explained by long-term memory model



Dynamic Causal Modelling

network structure identification



I thank you for your attention.

A note on statistical significance

lessons from the Neyman-Pearson lemma

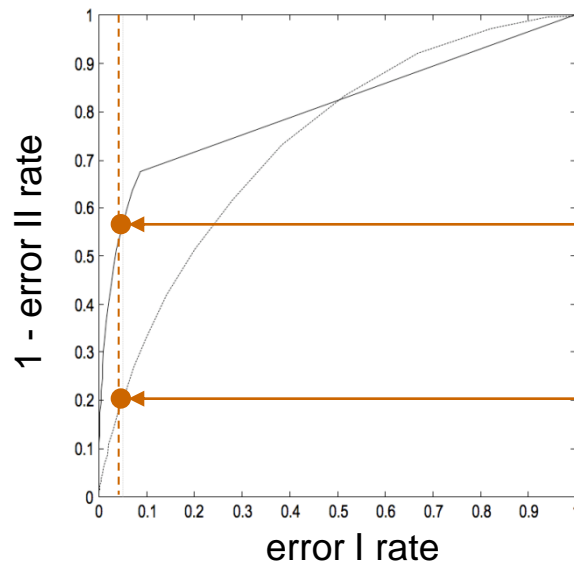
- **Neyman-Pearson lemma**: the likelihood ratio (or Bayes factor) test

$$\Lambda = \frac{p(y|H_1)}{p(y|H_0)} \geq u$$

is the most powerful test of size $\alpha = p(\Lambda \geq u | H_0)$ to test the null.

- what is the threshold u , above which the Bayes factor test yields a error I rate of 5%?

ROC analysis



MVB (Bayes factor)
 $u=1.09$, power=56%

CCA (F-statistics)
 $F=2.20$, power=20%